EM Algorithm for Maximum Likelihood Estimation of the Factor Model

Dylan Dijk

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Outline

Factor Model Recap

Factor Model Additional Assumptions

EM algorithm

EM algorithm Derivation for Factor Model

Appendix Derivation of EM steps for Factor Model

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Factor Models aim to explain the correlation between variables via a small number of k < p factors.

 $X = \mathbf{\Lambda}F + U$

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- $\bigwedge_{(p \times k)}$ is the **loadings matrix** of constants.
- *F* is an \mathbb{R}^k valued random variable, called the **factor**.

$$- \mathbb{E}[F] = 0 \quad Var(F) = I_k$$

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• U is an \mathbb{R}^p valued random variable

$$- \mathbb{E}[U] = 0 \quad Var(U) = \Psi = diag(\psi_{11}, \dots, \psi_{pp})$$

$$Cov(F, U) = 0$$

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- $Var(X) = \mathbf{\Lambda}\mathbf{\Lambda}^T + \mathbf{\Psi}$
- $Var(X|F) = \Psi$
- In lectures we used the **iterated principal factor analysis algorithm** to estimate Λ and Ψ from the correlation matrix R.

Additional Assumptions

If we make probabilistic assumptions for X and F we can then use MLE estimates.

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- $X|F \sim N_p(\Lambda F, \Psi)$
- $F \sim N_k(0, \boldsymbol{I}_k)$

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- $X|F \sim N_p(\Lambda F, \Psi)$
- $F \sim N_k(0, I_k)$

$$\implies \begin{pmatrix} F \\ X \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} I_k & \Lambda^T \\ \Lambda & \Lambda\Lambda^T + \Psi \end{pmatrix}\right)$$

 $\implies X \sim N(0, \mathbf{\Lambda}\mathbf{\Lambda}^T + \mathbf{\Psi})$

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Objective

$$X \sim N(0, \mathbf{\Lambda}\mathbf{\Lambda}^{T} + \mathbf{\Psi})$$

Given a dataset X, where rows are i.i.d copies of X, we want to estimate the parameters.

$$L(\mathbf{\Lambda}, \mathbf{\Psi}, \mathbf{X}) = \sum_{i=1}^{n} log(N(x_i; 0, \mathbf{\Lambda}\mathbf{\Lambda}^{T} + \mathbf{\Psi}))$$

No closed form solution for finding MLE estimates of Ψ and $\Lambda.$ Can use the EM algorithm.

- The EM algorithm is a very general technique for finding MLE solutions for probabilistic models with latent variables.
- Latent variables, Z, are variables that are not observed.
- EM algorithm is used in cases where direct optimisation of L(θ, X) := log(p(X; θ)) is difficult, but the optimisation of log(p(X, Z; θ)) is much easier.

EM Algorithm - Setup

- Complete-data likelihood $p(\mathbf{X}, \mathbf{Z}; \theta)$
- Incomplete-data likelihood $p(\mathbf{X}; \boldsymbol{\theta})$

We do not have the complete-data likelihood, so the idea is to maximise its expectation instead.

• E-Step

Compute the latent variable posteriors $p(z|x_i; \theta^{old})$

M-Step

$$\theta^{new} = \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^{n} \int_{z} p(z|x_{i}; \theta^{\text{old}}) \log p(x_{i}, z; \theta) dz$$
$$= \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^{n} \mathbb{E}_{Z|X=x_{i}; \theta^{\text{old}}} [log(p(x_{i}, z; \theta))]$$

The EM algorithm will increase the likelihood function $L(\theta, \mathbf{X})$

Let $q(\cdot)$ be a density over the latent variables Z. Then we can decompose the likelihood for a single observed value as:

 $L(\theta, x) = \mathcal{L}(q, \theta) + \operatorname{KL}(q||p)$

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$$\mathcal{L}(q,\theta) = \int_{z} q(z) \log \left\{ \frac{p(x,z;\theta)}{q(z)} \right\} dz$$
$$\mathsf{KL}(q||p) = -\int_{z} q(z) \log \left\{ \frac{p(z|x;\theta)}{q(z)} \right\} dz$$

$$L(\theta, x) = \mathcal{L}(q, \theta) + \operatorname{KL}(q||p)$$

 $\mathsf{KL}(q||p) \ge 0 \implies L(\theta, x) \ge \mathcal{L}(q, \theta)$



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• E-Step

Fixing a starting value of the parameters θ^{old} , $\mathcal{L}(q,\theta)$ is maximised with respect to q

$$\operatorname*{argmax}_{q} \mathcal{L}(q, \theta^{\mathsf{old}}) = p(z|x; \theta^{\mathsf{old}})$$

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$$L(\theta, x) = \mathcal{L}(q, \theta) + \operatorname{KL}(q||p)$$

$\mathsf{KL}(q||p) \ge 0 \implies L(\theta, x) \ge \mathcal{L}(q, \theta)$

• E-Step

Fixing a starting value of the parameters θ^{old} , $\mathcal{L}(q,\theta)$ is maximised with respect to q

$$\operatorname*{argmax}_{q} \mathcal{L}(q, \theta^{\mathsf{old}}) = p(z|x; \theta^{\mathsf{old}})$$

• M-Step

Now keeping q fixed and maximising with respect to θ

$$\underset{\theta}{\operatorname{argmax}} \mathcal{L}(q, \theta) = \underset{\theta}{\operatorname{argmax}} \mathbb{E}_{Z|X; \theta^{\text{old}}}[log(p(x, z; \theta))]$$

EM Steps for MLE of the Factor Model

Our model is constructed with a latent variable F, therefore we can use the EM algorithm with F in place of Z.

•
$$X|F \sim N_p(\Lambda F, \Psi)$$

•
$$F \sim N_k(0, \boldsymbol{I}_k)$$

$$\begin{pmatrix} F \\ X \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} I_k & \Lambda^T \\ \Lambda & \Psi + \Lambda \Lambda^T \end{pmatrix} \right)$$

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EM Steps for MLE of the Factor Model

• **E-Step** Compute $p(Z_i|X_i; \theta^{\text{old}})$

$$\mu_{F_i|X_i} = \mathbf{\Lambda}^T (\mathbf{\Psi} + \mathbf{\Lambda}\mathbf{\Lambda}^T)^{-1}(X_i)$$

$$\Sigma_{F_i|X_i} = \mathbf{I}_k - \mathbf{\Lambda}^T (\mathbf{\Psi} + \mathbf{\Lambda}\mathbf{\Lambda}^T)^{-1}\mathbf{\Lambda}$$

M-Step

$$\mathbf{\Lambda}^{\text{new}} = \left(\sum_{i=1}^{n} (X_i) \mathbb{E}[F_i]^T\right) \left(\sum_{i=1}^{n} \mathbb{E}[F_i F_i^T]\right)^{-1}$$
$$\Psi^{\text{new}} = \frac{1}{n} \text{diag} \left\{\sum_{i=1}^{n} \left(X_i X_i^T - \mathbf{\Lambda}^{\text{new}} \mathbb{E}[F_i] X_i\right)\right\}$$

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References

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 Bishop, Christopher (2006). Pattern Recognition and Machine Learning. URL:

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Derivation of EM steps for Factor Model

• **E-Step** Compute $p(Z_i|X_i; \theta^{\text{old}})$

Immediate from Gaussian identities

M-Step

$$\theta^{new} = \operatorname*{argmax}_{\theta} \sum_{i=1}^{n} \mathbb{E}_{Z_i | X_i; \theta^{\text{old}}}[log(p(X_i, Z_i; \theta))]$$

In our case we have that

$$p(X_i, F_i; \theta) = p(X_i | F_i; \theta) p(F_i)$$

 $p(F_i)$ does not depend on our parameters of interest.

Derivation of EM steps for Factor Model

$$Q := \sum_{i=1}^{n} \mathbb{E}_{F_i | X_i; \theta^{\text{old}}} [log(p(X_i | F_i; \theta))]$$

$$= \sum_{i=1}^{n} \mathbb{E} \left[\log \left((2\pi)^{p/2} |\Psi|^{-1/2} \exp \left\{ -\frac{1}{2} [X_i - \Lambda F_i]^T \Psi^{-1} [X_i - \Lambda F_i] \right\} \right) \right]$$

$$= c - \frac{n}{2} \log |\Psi|$$

$$- \sum_{i=1}^{n} \left(\frac{1}{2} X_i^T \Psi^{-1} X_i - X_i^T \Psi^{-1} \Lambda \mathbb{E}[F_i] + \frac{1}{2} \left[\mathbb{E}[F_i^T \Lambda^T \Psi^{-1} \Lambda F_i] \right] \right)$$

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M-Step

Now need to maximise with respect to $\theta = (\mathbf{\Lambda}, \mathbf{\Psi})$

$$\frac{\partial Q}{\partial \mathbf{\Lambda}} = \sum_{i=1}^{n} \underbrace{\frac{\partial}{\partial \mathbf{\Lambda}} X_{i}^{T} \Psi^{-1} \mathbf{\Lambda} \mathbb{E}[F_{i}]}_{(1)} - \frac{1}{2} \sum_{i=1}^{n} \mathbb{E} \left[\underbrace{\frac{\partial}{\partial \mathbf{\Lambda}} F_{i}^{T} \mathbf{\Lambda}^{T} \Psi^{-1} \mathbf{\Lambda} F_{i}}_{(2)} \right]$$

(1) =
$$\mathbb{E}[F_i](X_i^T \Psi^{-1})$$

(2) = $(2\Psi^{-1} \mathbf{\Lambda} F_i F_i^T)^T$

Setting to zero we get:

$$\boldsymbol{\Lambda}^{\mathsf{new}} = \left(\sum_{i=1}^{n} (X_i) \mathbb{E}[F_i]^T\right) \left(\sum_{i=1}^{n} \mathbb{E}[F_i F_i^T]\right)^{-1}$$

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M-Step

$$\frac{\partial Q}{\partial \boldsymbol{\Psi}^{-1}} = \frac{n}{2} \boldsymbol{\Psi}^{\text{new}} - \sum_{i=1}^{n} \left(\frac{1}{2} X_i X_i^T - \boldsymbol{\Lambda}^{\text{new}} \mathbf{E}[F_i] X_i^T + \frac{1}{2} \boldsymbol{\Lambda}^{\text{new}} \mathbf{E}[F_i F_i^T] (\boldsymbol{\Lambda}^{\text{new}})^T \right)$$

Setting to zero and plugging in $\pmb{\Lambda}^{new}$:

$$\frac{n}{2}\Psi^{\text{new}} = \sum_{i=1}^{n} \left(\frac{1}{2} X_i X_i^{\mathsf{T}} - \frac{1}{2} \mathbf{\Lambda}^{\text{new}} \mathbb{E}[F_i] X_i \right)$$

Restricting to a diagonal matrix Ψ :

$$\Psi^{\text{new}} = \frac{1}{n} \text{diag} \left\{ \sum_{i=1}^{n} \left(X_i X_i^T - \Lambda^{\text{new}} \mathbb{E}[F_i] X_i \right) \right\}$$